

### Some Poisson Random Variable Review Problems

Recall the following. Suppose a certain event happens, on average,  $\lambda$  times per interval of a given extent. Let  $X$  be the number of times that the event *actually* happens in such an interval. Then we say “ $X$  is Poisson of parameter  $\lambda$ ,” and we have the formula

$$P(X = k) = \frac{\lambda^k}{k!} e^{-\lambda} \quad (k = 0, 1, 2, 3, \dots). \quad (*)$$

1. On a given day, the Happa-One ski resort in Hakuba, Japan opens additional lifts depending on how many buses from Tokyo arrive at the resort by 11 AM. If fewer than 3 buses arrive by 11 AM, they open no additional lifts. If 3 or 4 buses arrive by 11 AM, they open one additional lift. If 5 or more arrive by 11 AM, they open two additional lifts.

Suppose, on average, 2 buses from Tokyo arrive at the resort by 11 AM.

- (a) Find the probability that, on a given day, Happa-One will open no additional lifts. Hint: this probability equals  $P(X = 0) + P(X = 1) + P(X = 2)$ , where  $X$  is the number of buses arriving by 11 from Tokyo.

$$\begin{aligned} P(\text{no additional lifts}) &= P(X = 0) + P(X = 1) + P(X = 2) \\ &= \frac{2^0}{0!} e^{-2} + \frac{2^1}{1!} e^{-2} + \frac{2^2}{2!} e^{-2} = 0.6767. \end{aligned}$$

- (b) Find the probability that, on a given day, Happa-One will open one additional lift.

$$\begin{aligned} P(\text{one additional lift}) &= P(X = 3) + P(X = 4) = \frac{2^3}{3!} e^{-2} + \frac{2^4}{4!} e^{-2} = 0.2707. \end{aligned}$$

- (c) Find the probability that, on a given day, Happa-One will open two additional lifts.

$$\begin{aligned} P(\text{two additional lifts}) &= 1 - (P(X = 0) + P(X = 1) + P(X = 2) + P(X = 3) + P(X = 4)) \\ &= 1 - (0.6767 + 0.2707) = 0.0526. \end{aligned}$$

- (d) Find the expected number of additional lifts that Happa-One will need to open on a given day. Call this number of lifts  $Y$ . Then

$$\begin{aligned} E[Y] &= 0 \cdot P(Y = 0) + 1 \cdot P(Y = 1) + 2 \cdot P(Y = 2) \\ &= 0 \cdot 0.6767 + 1 \cdot 0.2707 + 2 \cdot 0.0526 = 0.3759. \end{aligned}$$

2. A certain event happens, on average, 5 times per second.

(a) Find the probability that this event happens 9 times in a given second.

Let  $X$  be the number of times it happens in a second. Then

$$P(X = 9) = \frac{5^9}{9!} e^{-5} = 0.0363.$$

(b) Find the probability that it happens 18 times in a given two second interval.

If it happens, on average, 5 times a second, then on average, it happens, on average, 10 times every two seconds. So in this case,  $\lambda = 10$ . So if  $Y$  is the number of times it happens in a two second interval, Then

$$P(Y = 18) = \frac{10^{18}}{18!} e^{-10} = 0.0071.$$

(c) True or false: the event is just as likely to happen 18 times in two seconds as it is to happen 9 times in one second.

False, because  $P(Y = 18) = 0.0071$ , which is much less than half of half of  $P(X = 9) = 0.0363$ .

3. Show that the formula (\*) for the Poisson distribution, on page 1, really does define a probability distribution. This means: show that all the probabilities add up to one, or in other words,

$$\sum_{k=0}^{\infty} P(X = k) = 1.$$

Hint: you may want to use the Taylor series for  $e^\lambda$ :

$$e^\lambda = \sum_{k=0}^{\infty} \frac{\lambda^k}{k!}.$$

By (\*), we have

$$\sum_{k=0}^{\infty} P(X = k) = \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} e^{-\lambda}.$$

Since the factor  $e^{-\lambda}$  doesn't depend on  $k$ , we can pull it outside of the sum. We get

$$\sum_{k=0}^{\infty} P(X = k) = e^{-\lambda} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} = e^{-\lambda} e^\lambda,$$

by the power series for  $e^\lambda$ . But  $e^{-\lambda} e^\lambda = 1$ , so we're done.

4. In discussing the Poisson distribution in class, we first came up with the formula

$$P(X = k) = \lim_{N \rightarrow \infty} \binom{N}{k} \left(\frac{\lambda}{N}\right)^k \left(1 - \frac{\lambda}{N}\right)^{N-k}, \quad (**)$$

and then claimed that **(\*\*)** yields the formula **(\*)** on the first page of this review sheet. In class and in homework, we *proved* that **(\*\*)** yields **(\*)** for a few specific values of  $k$ . The purpose of this exercise is to show that this holds for *any* integer  $k = 0, 1, 2, \dots$

(a) Use the definition  $\binom{N}{k} = \frac{N!}{k!(N-k)!}$ , together with **(\*\*)**, to show that

$$P(X = k) = \frac{\lambda^k}{k!} \lim_{N \rightarrow \infty} \frac{N!}{N^k (N-k)!} \left(1 - \frac{\lambda}{N}\right)^{N-k}.$$

By **(\*\*)**, we have

$$\begin{aligned} P(X = k) &= \lim_{N \rightarrow \infty} \binom{N}{k} \left(\frac{\lambda}{N}\right)^k \left(1 - \frac{\lambda}{N}\right)^{N-k} \\ &= \lim_{N \rightarrow \infty} \frac{N!}{k!(N-k)!} \left(\frac{\lambda}{N}\right)^k \left(1 - \frac{\lambda}{N}\right)^{N-k}. \end{aligned}$$

The factor  $\lambda^k$  in the numerator and the factor  $k!$  in the denominator are independent of  $N$ , so we can bring them outside of the limit, to get

$$\begin{aligned} P(X = k) &= \frac{\lambda^k}{k!} \lim_{N \rightarrow \infty} \frac{N!}{(N-k)!} \left(\frac{1}{N}\right)^k \left(1 - \frac{\lambda}{N}\right)^{N-k} \\ &= \frac{\lambda^k}{k!} \lim_{N \rightarrow \infty} \frac{N!}{N^k (N-k)!} \left(1 - \frac{\lambda}{N}\right)^{N-k}. \end{aligned}$$

(b) Use your answer to the previous part of this problem to find a simple formula (with no limits in it) for  $P(X = k)$ . Hint: use the following limit formulas (which may be proved using calculus ideas, like l'Hôpital's rule):

$$\lim_{N \rightarrow \infty} \frac{N!}{N^k (N-k)!} = 1; \quad \lim_{N \rightarrow \infty} \left(1 - \frac{\lambda}{N}\right)^N = e^{-\lambda}, \quad \lim_{N \rightarrow \infty} \left(1 - \frac{\lambda}{N}\right)^{-k} = 1.$$

Another hint: you know what your answer should be, by **(\*)**.

By our previous answer and by the given limits, we have

$$\begin{aligned} P(X = k) &= \frac{\lambda^k}{k!} \lim_{N \rightarrow \infty} \frac{N!}{N^k (N-k)!} \left(1 - \frac{\lambda}{N}\right)^{N-k} \\ &= \frac{\lambda^k}{k!} \cdot 1 \cdot e^{-\lambda} \cdot 1 = \frac{\lambda^k}{k!} e^{-\lambda}, \end{aligned}$$

which agrees with **(\*)**.