

Introduction to Probability
HW3 Solutions

(3.1) Directly from the definition of conditional probability:

$$\begin{aligned} P(\text{at least one six} \mid \text{each roll different}) &= \frac{P(\text{one six and the other not})}{P(\text{each roll different})} \\ &= \frac{10/36}{30/36} = 1/3. \end{aligned}$$

(3.5) Condition to find

$$\begin{aligned} P(\circ \circ \bullet \bullet) &= P(\circ \circ \bullet \bullet \mid \text{pick 2 } \circ \text{'s and 2 } \bullet \text{'s})P(\text{pick 2 } \circ \text{'s and 2 } \bullet \text{'s}) \\ &= \frac{1}{\binom{4}{2}} \cdot \frac{\binom{6}{2}\binom{9}{2}}{\binom{15}{4}} = 6/91. \end{aligned}$$

(3.24) Point is to understand what the conditional info is telling you. To start there are 4 scenarios: GB, BG, GG, BB . Noting that the first two are different is important.

a) Knowing that gold paint was used rules out BB . Therefore, the desired event is now one of three possibilities and the conditional probability is $1/3$.

b) This is a question about just one ball, and for either we have that $P(G) = 1/2$.

(3.28) Using that the conditional space is all lists of 52 cards where we see no ace until card (or place) number twenty we compute:

a) $P(\text{card twenty-one is Ace of Spades} \mid \text{first ace on card 20})$

$$= \frac{48 \cdot 47 \cdots 30 \cdot 3 \cdot 1 \cdot 31 \cdot 30 \cdots 1}{48 \cdot 47 \cdots 30 \cdot 4 \cdot 32 \cdot 31 \cdot 30 \cdots 1} = \frac{3}{4 \cdot 32}.$$

b) $P(\text{card twenty-one is Two of Clubs} \mid \text{first ace on card 20})$

$$= \frac{47 \cdot 46 \cdots 29 \cdot 4 \cdot 1 \cdot 31 \cdot 30 \cdots 1}{48 \cdot 47 \cdots 30 \cdot 4 \cdot 32 \cdot 31 \cdot 30 \cdots 1} = \frac{29}{48 \cdot 32}$$

as now both aces and the two of clubs (that is, five cards) must be avoided in the first nineteen places comprising the numerator.

(3.38) This is “straight-up” Bayes:

$$\begin{aligned} P(\text{coin is T} \mid \text{see White}) &= P(\text{used urn B} \mid W) \\ &= \frac{P(W \mid \text{urn B})P(\text{urn B})}{P(W \mid \text{urn B})P(\text{urn B}) + P(W \mid \text{urn A})P(\text{urn A})} \\ &= \frac{3/15}{5/12 + 3/15} = 12/37. \end{aligned}$$

(3.57) A question in independence:

a) $P(\text{no move after 2 days}) = P(UD) + P(DU) = 2P(U) \cdot P(D) = 2p(1-p)$.

b) $P(\text{up one after 3 days}) = P(UUD) + P(UDU) + P(DUU) = 3p^2(1-p)$

c) $P(\text{day one is up} \mid \text{up one after 3 days}) =$

$$\frac{P(UDU) + P(UUD)}{P(UDU) + P(UUD) + P(DUU)} = 2/3.$$

(TH 3.6) The real point is to realize that if $\{E_1, E_2, \dots, E_n\}$ are all independent then so are $\{E_1^c, E_2^c, \dots, E_n^c\}$ and therefore

$$\begin{aligned} P\left(\bigcup_{i=1}^n E_i\right) &= 1 - P\left(\bigcap_{i=1}^n E_i^c\right) \\ &= 1 - \prod_{i=1}^n P(E_i^c) = 1 - \prod_{i=1}^n (1 - P(E_i)) \end{aligned}$$

where DeMorgan is used in line one.

(TH 3.7 (a)) The hint is to think about lining all the balls in a row (after removing them all, where you line them up in the order that they were removed). Then the question becomes: what is the probability that the last several (however many) are white. This turns out to be the same as asking what is the probability the last is white:

$$\begin{aligned} &P(\text{last few are white}) \\ &= P(\text{last few are white} \cap \text{last is white}) + P(\text{last few are white} \cap \text{last is black}) \\ &= P(\text{last few are white} \mid \text{last is white}) \cdot P(\text{last is white}) + 0 \\ &= 1 \cdot P(\text{last is white}) + 0 \\ &= \frac{n}{n+m}. \end{aligned}$$